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# Relationships between poverty and AIDS Illness in South Africa: an investigation of urban and rural households in KwaZulu-Natal

Janina Isabel Steinert, Lucie Cluver, G.J. Melendez-Torres<sup>†</sup> and Rocio Herrero Romero

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## ABSTRACT

The association between poverty and HIV/AIDS in Sub-Saharan Africa remains contested. A better understanding of the relationship between the prevalence of poverty and the disease is essential for addressing prevention, treatment, and care. The present study interrogates this relationship, using a cross-sectional survey of 2477 households in urban and rural KwaZulu-Natal, South Africa. Structural equation modelling was employed to estimate the correlations between poverty and AIDS illness. The analysis revealed a correlation of  $r_{pb} = 0.23$ , denoting that a higher level of household poverty was associated with a higher likelihood of being AIDS-unwell. *Post hoc t*-test showed that receipt of a disability grant by AIDS-affected households was associated with significantly lower poverty, compared to AIDS-affected households not receiving the grant,  $t(654) = 3.67$ ,  $p < .01$ . Geographic location was found to confound the correlation: the strength of the relationship between poverty and AIDS was decreased to  $r_{pb} = 0.15$  ( $p < .001$ ) for the urban and  $r_{pb} = 0.16$  ( $p < .001$ ) for the rural sub-population. Findings suggest the importance of two sets of policies: those that address the potential upstream risk of poverty through economic interventions, and those that alleviate the impoverishing effects of AIDS illness for affected households.

## ARTICLE HISTORY

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## KEYWORDS

HIV/AIDS; poverty; structural equation modelling; poverty measurement; verbal autopsy


## Introduction

Globally, 35 million people are living with HIV/AIDS (UNAIDS, 2013). Around 95% of them are localised in low- and middle-income countries, leading to claims that poverty and HIV/AIDS go hand in hand (Lachaud, 2007). However, the nature of this association has caused substantial debate in the literature over the last years (Butler, 2000; Hargreaves, Davey, & White, 2012).

Numerous studies have found a positive correlation between HIV/AIDS and poverty (Chakraborty, Firestone, & Bellows, 2013; Cluver, Orkin, Boyes, Gardner, & Meinck, 2011; Fenton, 2004; Madise, Zulu, & Ciera, 2007; Masanjala, 2007; Shisana, Rice, Zungu, & Zuma, 2010; UNAIDS, 2001). This correlation is likely to reflect both a hypothesised *upstream effect* of poverty on the likelihood of HIV infection as well as

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a *downstream effect* of AIDS illness on households' poverty levels. The *upstream hypothesis* argues that impoverished people – largely women – are more likely to contract HIV due to increased necessity for transactional and age-disparate sexual relationships (Krishnan et al., 2008; Mojola, 2014; Wojcicki, 2005). Lack of education and limited HIV-related knowledge are other causal pathways (Bärnighausen, Hosegood, Timaeus, & Newell, 2007; deWalque, Nakiyingi-Miir, Busingye, & Whitworth, 2005). Furthermore, chronic malnutrition, inadequate sanitation, and limited access to health infrastructure can increase susceptibility to HIV infection among people living in poverty (Poku, 2002; Solomon, 2012; Wojcicki, 2005). Vice versa, in accordance with the *downstream hypothesis*, studies suggest that AIDS illness can exacerbate the level of poverty in affected households/families (Bachmann & Booysen, 2003, 2006; Booysen & Arntz, 2003; Salinas & Haacker, 2006). The economic burden of an AIDS-unwell household member amplifies in direct costs such as medical expenses and costs for transport to health care centres (Bollinger & Stover, 1999; McIntyre, Thiede, Dahlgren, & Whitehead, 2006; Russell, 2004; Whiteside, 2002). Further indirect costs may include loss of income (particularly when the primary breadwinner is ill), diversion from productive labour to caring for the sick, and removal of children from school and thus the inter-generational transmission of poverty (Booyesen, 2004; Loewenson, Hadingham, & Whiteside, 2009; Mahal, Canning, Odumosu, & Okonkwo, 2008; Poku, 2002).

Conversely, some studies suggest an opposite relationship: that household wealth rather than poverty is linked to HIV/AIDS (Chin, 2007; Fox, 2010, 2012; Fortson, 2008; Gillespie, Kadiyala, & Greener, 2007; Mishra et al., 2007, 2008; Parkhurst, 2010; Shelton, Cassell, & Adetunji, 2005). On a country level, higher prevalence rates are observed in the wealthier countries of Sub-Saharan Africa such as South Africa. On an individual level, relative wealth is associated with greater mobility and access to greater social networks, thus facilitating multiple sexual partnerships (Fox, 2012; Nattrass, 2009; Parkhurst, 2010).

In light of this debate, it is imperative to further investigate the relationship between AIDS illness and poverty. While great strides have been made in the global fight against HIV/AIDS and antiretroviral therapy has been rolled out in the developing world, it is essential to examine whether AIDS illness and poverty remain interlinked. The present paper therefore sets out to test whether there is a positive correlation between poverty and AIDS illness and whether the correlation holds across geographic locations within South Africa. Understanding of the question at hand is important for HIV prevention efforts as well as current considerations of 'HIV-sensitive social protection' provision (UNICEF, 2012).

## Methods

### Participants

This study used cross-sectional data that were collected by the authors for a study on child abuse, HIV/AIDS, and poverty. Data collection took place between 2009 and 2010 in the KwaZulu-Natal province in both rural (Manguzi/kwaNgwanase) and urban (Lamontville townships) communities, sampling 2477 households. Survey respondents were recruited through household stratified systematic random sampling. Census enumeration areas and designated tribal authority areas were identified via Geographical Information System and respondents were selected with door-to-door household sampling.

Households were eligible for participation in the study if they had a resident child/adolescent between the ages of 10 and 17. One child/adolescent per household was randomly selected and asked to identify a primary caregiver, defined as the person staying with and caring for the child – potentially biological parents, other family members, or non-relatives. KwaZulu-Natal records the highest HIV prevalence in South Africa (>35%) and more than two-thirds of its population are living on less than US\$2 per day (Statistics South Africa, 2012; South African National Department of Health, 2011). Despite a claim for universal access to antiretroviral treatment (ART), the roll-out of ART in KwaZulu-Natal has been hindered by insufficiencies in health infrastructure, particularly so in rural areas (Michel & Matlakala, 2013; Ojikutu, 2007). In addition, food insecurity, lack of education, and the stigma associated with HIV/AIDS pose a serious threat to good ART adherence (Michel, Matlakala, English, Lessells, & Newell, 2013). AIDS symptomatology as well as mortality thus remain high.

### Procedures

Trained and supervised local research assistants conducted 40-to-60-minute structured face-to-face interviews with standardised questionnaires. No incentives were provided other than refreshments and certificates of appreciation. Participation was voluntary. Ethical protocols were approved by the University of Oxford, University of KwaZulu-Natal, and the provincial Departments of Health and Education.

### Measurements

#### AIDS illness

The analysis considered primary caregivers to be AIDS-unwell if they (1) indicated use of antiretrovirals, (2) had a CD4 count <200 cells/ $\mu$ L, or (3) met a conservative threshold of three or more *current* AIDS-related symptoms from the adapted verbal autopsy (VA) checklist, such as severe weight loss, oral candidiasis, or Kaposi's sarcoma (see Table A in supplemental data). The VA method was developed by Lopman et al. (2006) for post-mortem classification of AIDS-associated deaths in areas with HIV prevalence rates above 20%. Recent studies conducted in South Africa have validated the VA with 89% sensitivity, 93% specificity, and a positive prediction rate of 76% (Hosegood, Vanneste, & Timaeus, 2004; Kahn, Tollman, Garenne, & Gear, 2000). There has been an ongoing debate on when to classify a person as having AIDS, particularly in the light of treatment and recovery from severe illness. However, given that both AIDS illness as well as being on ART likely have implications for poverty, this paper uses the term 'AIDS-unwell' to refer to both current AIDS illness as well as a history of AIDS (i.e. a CD4 count of <200 cells/ $\mu$ L) resulting in ART initiation. While the VA method was designed to reduce the stigmatisation and shame that are associated with AIDS illness (Kalichman & Simbayi, 2003), it does not allow for identifying asymptomatic HIV-positive individuals.

#### Poverty

Conventional approaches to the measurement of poverty that rely on income or expenditure data can have limitations in low-income contexts. Income data are highly susceptible to measurement errors through seasonal fluctuations in earnings and informal borrowing

mechanisms (Moser & Felton, 2007; Sahn & Stifel, 2003). Collecting expenditure/consumption data can be time-consuming and potentially subject to recall bias (Béranger, Deutsch, & Silber, 2013; Klasen, 2000; Sahn & Stifel, 2003). Finally, it may be valuable to capture the multidimensional character of poverty (Batana, 2013; Béranger et al., 2013; OPHI, 2013). This study thus constructs a poverty scale that is based on the various household-level deprivations (Alkire & Foster, 2011; Carter & Barrett, 2006; Filmer & Scott, 2012; Harttgen & Vollmer, 2013; Moser & Felton, 2007; Wright & Noble, 2007). The poverty indicators presented in Table 1 were adopted from the South African *Demographic and Health Survey* and draw on previous work on composite poverty measurements (see Filmer & Pritchett, 2001; Sahn & Stifel, 2003). While we use the term poverty throughout this paper, the constructed measurement is likely to capture other related concepts and terms such as household economic well-being or material and human deprivation.

## Analysis

In the first step, traditional exploratory factor analysis (EFA) was employed to explore which indicators from Table 1 can adequately measure poverty (see Acock, 2013). EFA was based on a polychoric correlation matrix to accommodate the binary or ordinal nature of the indicators (Holgado-Tello, Chacón-Moscoso, Barbero-García, & Vila-Abad, 2010). Scree plots were inspected for choosing principal components. Following Costello and Osborne (2005), a factor loading of 0.3 and above was deemed the minimum criterion for an item.

Subsequently, structural equation modelling (SEM) was introduced to estimate the point-biserial correlation between poverty and HIV/AIDS. Maximum likelihood estimators were used to fit the suggested models (Acock, 2013; StataCorp, 2013).

Following this, control variables were added to the model: *sex* and *age of the primary caregiver*, *urban/rural residency*. In addition, the analysis accounted for the *number of adults living in the household* (aged between 25 and 49 years) which may be linked to a family's potential for income generation as well as shared caregiving. Lastly, a binary variable as of whether the primary caregiver was receiving a *disability grant* at the time of measurement was added with the hypothesis that access to grants might alleviate the economic burden of AIDS illness for households. While families in South Africa may have access to a variety of

**Table 1.** Poverty indicators used for the analysis.

Housing quality	Asset ownership	Human capital
<ul style="list-style-type: none"> <li>• Source of water</li> <li>• Safe drinking water</li> <li>• Electricity</li> <li>• Cooking fuel</li> <li>• Heating fuel</li> <li>• Lighting</li> <li>• Toilet facilities</li> <li>• Dwelling type</li> <li>• Wall type</li> <li>• Floor type</li> <li>• Number of rooms/overcrowding</li> </ul>	<ul style="list-style-type: none"> <li>• Transport (bicycle, motorcycle/car)</li> <li>• Refrigerator</li> <li>• Washing machine</li> <li>• TV</li> <li>• Computer</li> <li>• Telephone</li> <li>• Radio</li> <li>• Food/hunger</li> <li>• Meat once a week</li> <li>• Livestock (horse, donkey, cattle, sheep)</li> </ul>	<ul style="list-style-type: none"> <li>• Education (of primary caregiver)</li> <li>• Employment (of any household member)</li> <li>• Schooling (of any household member)</li> </ul>

welfare grants, eligibility for the disability grant is defined through physical/mental impairment and thus can be received by some AIDS-unwell individuals (SASSA, 2010).

Lastly, a multiple-group comparison was performed in order to examine whether correlation estimates differ for specific subgroups. Focus was placed on urban/rural residency which was previously found to be a potential confounder in the relationship between poverty and HIV/AIDS (Bachmann & Booyesen, 2003; Fox, 2010, 2012; Wojcicki, 2005). For this analysis, the model with the least constraints (configural invariance) was chosen to allow differential factor loadings for individual poverty indicators across areas (see Steenkamp & Baumgartner, 1998). In so doing, we seek to build more adequate poverty measures for the specific population of interest.

Results

Sociodemographic characteristics of the population sample are summarised in Table 2. Across all 2477 households, 27.4% of the interviewed caregivers were identified as AIDS-unwell according to the adapted VA method. For all indicators, ownership or achievement was higher in households with healthy primary caregivers, significantly for most items (see Figure 1).

EFA

EFA of the poverty indicators showed five factors (see Table B in supplemental data) with Eigenvalues greater than 1 (see Kaiser, 1974) and the first factor explaining 44.7% of the variance (see Figure 2). However, after varimax rotation, items on each factor failed to correspond meaningfully to any putative dimensions of poverty such as asset ownership, housing quality, or human capital. We therefore selected a single-factor solution (Table 3). In addition, EFA was displayed for the urban and rural

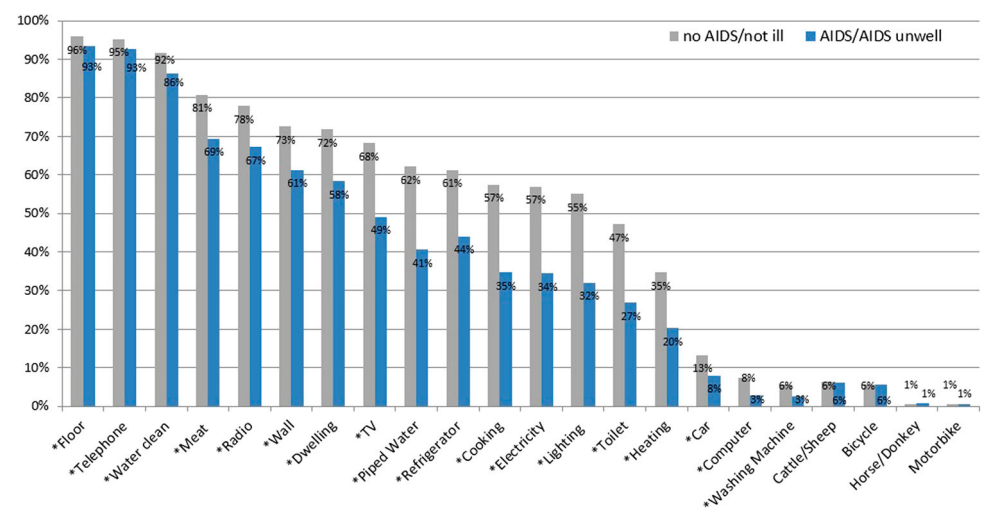


Figure 1. Binary Poverty Indicators by AIDS status of the primary caregiver. Note: items marked with \* show significant differences between households with AIDS-unwell and households with 'healthy' caregivers.

**Table 2.** Sample characteristics.

<i>Binary/categorical variables</i>	Complete sample <i>n</i> = 2477 <i>n</i> (%)	No AIDS illness <i>n</i> = 1798 (72.6%) <i>n</i> (%)	AIDS illness <i>n</i> = 679 (27.4%) <i>n</i> (%)	Difference <i>p</i> -value
On antiretroviral therapy	272 (11.0%)	0 (0%)	272 (40.1%)	
<i>Residency</i>				
Rural	1279 (51.7%)	825 (45.9%)	454 (66.9%)	<i>p</i> < .001
Urban	1198 (48.3%)	973 (54.1%)	225 (33.1%)	
<i>Gender</i>				
Female	2199 (88.8%)	1597 (88.9%)	602 (88.7%)	n.s.
Male	278 (11.2)	201 (11.1%)	77 (11.3%)	
<i>Population group</i>				
African/Black	2473 (99.9%)	1795 (99.8%)	678 (99.9%)	n.s.
Coloured	3 (0.1%)	2 (0.1%)	1 (0.1%)	
White	1 (<0.1%)	1 (<0.1%)		
<i>Education</i>				
No schooling	493 (20.1%)	315 (17.6%)	178 (26.5%)	<i>p</i> < .001
Primary	661 (26.9%)	447 (25.0%)	214 (31.8%)	
Secondary	871 (35.4%)	657 (36.8%)	214 (31.8%)	
Matric	407 (16.5%)	343 (19.2%)	64 (9.5%)	
University	28 (1.1%)	25 (1.4%)	3 (0.4%)	
<i>Employment</i>				
Permanent job	304 (12.4%)	246 (13.8%)	58 (8.7%)	<i>p</i> < .001
Temporary job	351 (14.3%)	264 (14.8%)	87 (13.0%)	
Unemployed	1794 (73.3%)	1271 (71.4%)	523 (78.3%)	
<i>Main source of income</i>				
Salaries and/or wages	868 (35.1%)	710 (39.6%)	158 (23.3%)	<i>p</i> < .001
Remittances	61 (2.5%)	48 (2.7%)	13 (1.9%)	
Welfare grants <sup>a</sup>	1255 (50.7%)	835 (46.5%)	420 (61.9%)	
Sales of farm products/services	84 (3.4%)	58 (3.2%)	26 (3.8%)	
No income	94 (3.8%)	66 (3.7%)	28 (4.1%)	
Other	112 (4.5%)	78 (4.3%)	34 (5%)	
<i>Disability grant</i>				
Household receipt of grant	160 (6.5%)	90 (5.0%)	70 (10.3%)	<i>p</i> < .001
<i>Continuous variables</i>	<i>M</i> ( <i>SD</i> )	<i>M</i> ( <i>SD</i> )	<i>M</i> ( <i>SD</i> )	
Number of AIDS-related symptoms	1.39 (1.90)	0 (0)	3.75 (2.04)	
Caregiver age	44.22 (13.88)	44.92 (14.60)	42.33 (11.57)	<i>p</i> < .001
Household size	6 (2.67)	6.04 (2.71)	5.88 (2.55)	n.s.
Number of adults (25–49 years)	2.07 (1.17)	2.14 (1.20)	1.90 (1.05)	<i>p</i> < .001
Children not in school (average)	0.79 (1.46)	0.77 (1.47)	0.84 (1.45)	n.s.
Children going to school (average)	2.34 (1.31)	2.41 (1.34)	2.31 (1.30)	n.s.

Note: n.s. = not significant, missing values for education, employment, main source of income.

<sup>a</sup>Including: Child Support Grant for the alleviation of child poverty, Foster Care Grant, Care Dependency Grant, Old Age Pension, and Government Housing Subsidy.

Pearson's chi-square test was applied to test differences between households with and without AIDS-unwell caregivers for binary/ordinal outcome variables, and an independent *t*-test was performed for continuous outcome variables.

sample separately. As results imply some differences in factor loadings between locations, we only removed items where loadings were below 0.3 in both urban and rural locations.

### Structural equation model

Goodness of fit of the original poverty measurement model proved low. Therefore, modification indices were inspected and errors for some item pairs were correlated accordingly (see Table C in supplemental data).

The above procedure generated a continuous poverty scale centred around zero (standardised scale: mean  $-0.001$ , 95% CI  $[-0.007, 0.004]$ , min  $-0.17$ , max  $0.16$ ), whereby



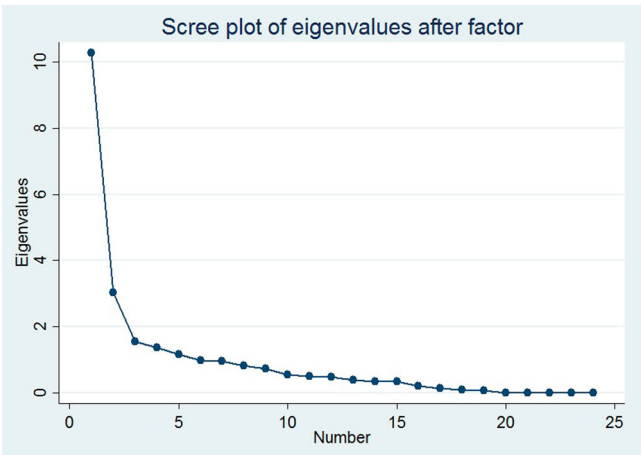


Figure 2. Scree plot: EFA.

larger scale scores denote higher levels of poverty (or higher levels of human and material deprivation).

Table 4 summarises the results from SEM analyses. The point-biserial correlation was  $r_{pb} = 0.22$  ( $p < .001$ ), implying a substantial association between poverty and AIDS illness. Covariates were introduced subsequently, but the correlation coefficient remained virtually unchanged. As depicted in Figure 3, households with AIDS-unwell primary caregivers were ranked on average as poorer than those with ‘healthy’ primary caregivers. A

Table 3. Summary of single-factor EFA (by location).

Item	Full sample (N = 2352)	Rural sample (n = 1212)	Urban sample (n = 1140)
Hunger	0.25	0.36	0.21
Refrigerator	−0.74	−0.59	−0.64
Phone	−0.36	−0.15	−0.18
Computer	−0.64	−0.20	−0.24
Transport	−0.36	−0.44	−0.28
Meat	−0.63	−0.36	−0.25
TV	−0.87	−0.55	−0.68
Radio	−0.65	−0.36	−0.44
Drinking source	−0.85	−0.31	−0.58
Safe water	−0.54	−0.01	−0.13
Washing machine	−0.07	−0.11	−0.22
Electricity	−0.74	−0.58	−0.72
Cooking	−0.94	−0.54	−0.16
Heating	−0.91	−0.39	−0.30
Lighting	−0.85	−0.61	−0.68
Toilet	−0.93	−0.24	−0.41
Floor	−0.46	−0.25	−0.20
Wall	−0.46	−0.54	−0.68
Dwelling	−0.55	−0.55	−0.66
Overcrowding	0.07	0.30	0.02
Livestock	0.28	−0.14	−0.07
Education	−0.67	−0.36	−0.21
Employment	−0.32	−0.25	−0.10
Schooling	−0.22	−0.06	−0.01
Eigenvalues	6.47	3.56	4.01
% of variance	75.6%	67.5%	60.8%
$\alpha$	0.86	0.70	0.72



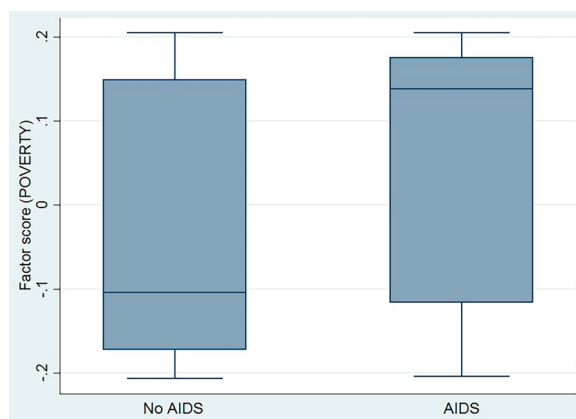
standardised-mean difference was calculated and yielded a *Cohen's d* of 0.52 (95% CI [0.43, 0.61]). In a similar vein, the odds of being identified as AIDS-unwell according to the VA method were 1.73 times higher (SE = 0.086) with a one-unit increase on the poverty scale.

Receipt of the disability grant was more prevalent in households with an AIDS-unwell caregiver and the level of poverty was significantly and negatively correlated with access to the disability grant. Looking at households with AIDS-unwell caregivers in specific, a *post hoc t*-test revealed that households that received the grant had a significantly lower average score on the standardised poverty scale,  $t(654) = 3.67$ ,  $p < .01$ . The mean score on the composite poverty scale was 0.03 (95% CI [-0.003; 0.05]) in AIDS-affected households receiving the grant in comparison to a mean score of 0.06 (95% CI [0.05; 0.07]) in AIDS-affected households not receiving the grant, thus indicating a higher level of poverty in the latter.

The analysis revealed a substantial correlation between poverty and rural location ( $r = 0.97$ ,  $p < .001$ ) and points to a potential risk of multicollinearity. It is therefore crucial to proceed with estimating the relationship between poverty and AIDS for urban/rural sub-groups separately in order to filter out the confounding effects of residency.

### Multiple-group comparison

Table 5 presents the results from multiple-group comparisons in SEM, stratifying the model by urban and rural location. Most importantly, the correlation between poverty and AIDS illness was diminished. Coefficients remained positive and significant ( $p < .001$ ), but were only  $r_{pb} = 0.15$  for the urban and  $r_{pb} = 0.16$  for the rural sub-population. The model was defined to allow for differential factor loadings for individual poverty indicators between rural and urban datasets. As the results show, some factor loadings differed substantially between locations. This finding likely indicates that the experience of poverty differs between urban and rural households and that the importance of specific poverty items varies across locations. As a robustness check, two fully separate models were run, stratified by urban and rural location, but differences in results were negligibly minor.



**Figure 3.** Boxplots: poverty by AIDS status of the primary caregiver.

**Table 4.** Final SEM results for the effect of AIDS illness on poverty.

	AIDS illness and poverty
<i>Structural model</i>	
Correlation of AIDS illness and poverty	$r_{pb} = 0.22^{***} (0.020)$
<i>Measurement model</i>	
Hunger	0.16***
Meat	−0.36***
Computer	−0.20***
TV	−0.61***
Phone	−0.10***
Radio	−0.37***
Refrigerator	−0.65***
Drinking source	−0.67***
Safe water	−0.28***
Washing machine	−0.24***
Electricity	−0.89***
Cooking	−0.95***
Heating	−0.69***
Lighting	−0.92***
Toilet	−0.85***
Floor	−0.11***
Wall	−0.13***
Dwelling	−0.21***
Overcrowding	−0.04*
Transport	−0.06***
Education	−0.59***
Employment	−0.19***
<i>Covariances</i>	
AIDS illness with gender	−0.00
AIDS illness with age	−0.08***
AIDS illness with grant	0.10***
AIDS illness with rural	0.19***
AIDS illness with the number of adults	−0.09***
Poverty with gender	0.05*
Poverty with age	−0.07**
Poverty with grant	−0.05*
Poverty with rural	0.96***
Poverty with number of adults	−0.20***
Error × meat with error × hunger	−0.26***
Error × electricity with error × lighting	0.67***
Error × floor with error × dwelling	0.06***
Error × wall with error × dwelling	0.73***
Error × education with error × employment	0.15**
<i>Goodness of fit</i>	
$\chi^2$	4820.829***
CFI	0.85
RMSEA	0.08
SRMR	0.07
N	2357

Note: Standard errors in parentheses.

\* $p < .05$ .  
 \*\* $p < .01$ .  
 \*\*\* $p < .001$ .

### Discussion

The present analysis found evidence of a significant and positive relationship between poverty and HIV/AIDS on a household level, in contrast to claims that wealth is a predictor of HIV in Sub-Saharan African countries. It is important to highlight that the sample of this study was drawn from economically deprived communities and was thus

**Table 5.** Multiple-group SEM for urban and rural sub-populations.

	Urban (N = 1144)		Rural (N = 1212)	
	Standardised	Unstandardised	Standardised	Unstandardised
<i>Structural model</i>				
Correlation of AIDS illness and poverty	$r_{pb} = 0.15^{***}$	$B = 0.01^{***}$	$r_{pb} = 0.16^{***}$	$B = 0.02^{***}$
<i>Measurement model</i>				
Hunger	0.21***	1 (fixed)	0.36***	1 (fixed)
Meat	-0.26***	-0.44***	-0.34***	-0.52***
Computer	-0.20***	-0.35***	-0.21***	-0.10***
TV	-0.76***	-1.37***	-0.57***	-0.91***
Phone	-0.22***	-0.24***	-0.15***	-0.13***
Radio	-0.48***	-0.90***	-0.37***	-0.59***
Refrigerator	-0.71***	-1.47***	-0.61***	-0.91***
Drinking source	-0.58***	-1.11***	-0.32***	-0.46***
Safe water	-0.14***	-0.10***	-0.02	-0.03
Washing machine	-0.18***	-0.32***	-0.10***	-0.02***
Electricity	-0.65***	-0.86***	-0.52***	-0.50***
Cooking	-0.16***	-0.08***	-0.54***	-0.42***
Heating	-0.29***	-0.84***	-0.40***	-0.17***
Lighting	-0.60***	-0.83***	-0.55***	-0.44***
Toilet	-0.32***	-0.68***	-0.24***	-0.11***
Floor	-0.17***	-0.16***	-0.23***	-0.18***
Wall	-0.55***	-1.45***	-0.42***	-0.63***
Dwelling	-0.52***	-1.33***	-0.44***	-0.67***
Education	-0.18***	-0.79***	-0.38***	-1.15***
Employment	-0.09***	-0.41***	-0.24***	-0.46***
Overcrowding	-0.01	-0.35	-0.28***	-1.26***
Transport	-0.26***	-0.51***	-0.46***	-0.56***
<i>Covariances</i>				
Error × meat with error × hunger	-0.16***	-0.04***	-0.22***	-0.08***
Error × electricity with error × lighting	0.83***	0.03***	0.58***	0.03***
Error × floor with error × dwelling	0.04	0.01	0.06*	0.01*
Error × wall with error × dwelling	0.69***	0.10***	0.64***	0.12***
Error × education with error × employment	0.19***	0.11***	0.07*	0.03*
<i>Goodness of fit</i>				
$\chi^2$	2721.5***			
CFI	0.80			
RMSEA	0.07			
SRMR	0.06			

\* $p < .05$ .\*\* $p < .01$ .\*\*\* $p < .001$ .

The model was fitted under the assumption of configural invariance.

impoverished *per se*. Yet, by adopting a wider conceptualisation of household poverty (that goes beyond household income), this empirical analysis can provide a nuanced understanding of the links between AIDS unwellness and experienced economic deprivation. The strength of the association was considerably reduced when accounting for urban/rural residency in multiple-group comparisons. That is, the initial correlation between poverty and AIDS was likely inflated both by the strong association between rural residency and the composite poverty measurement as well as by the link between rural residency and higher AIDS rates. Failure to account for rural and urban household status would thus over-estimate effect sizes. Notwithstanding, the correlation coefficient in the current study remained positive and significant for both subgroups. Therefore, the findings suggest that poverty and HIV/AIDS are interlinked, but concur with wider literature suggesting that the association is multi-faceted and contextual and that other factors such as labour migration, religion, and the historic legacy of a country can alter this association

(Auvert et al., 2001; Gizelis, 2009; Kalichman et al., 2006; Weiss & McMichael, 2004). Likewise, household poverty is driven by multiple factors including rural underdevelopment (Adato, Carter, & May, 2006; Woolard & Klasen, 2004).

Some tentative hypotheses can be formulated based on the above results. A higher number of adult household members were significantly associated with a lower level of poverty and a reduced prevalence of AIDS-unwell caregivers. Hence, single-headed households might be less resilient to AIDS-induced economic shocks as no other adult in the household can guarantee income generation and caregiving.

Further, the majority of AIDS-unwell caregiver in this sample indicated government grants as the main source of income of the household. This is in line with previous studies showing that income losses resulting from AIDS illness are partly substituted with welfare payments in countries that provide them (see Murtin & Marzo, 2013; Raniga & Simpson, 2011; Schatz, Madhavan, & Williams, 2010; Tladi, 2006). For instance, qualitative case studies in KwaZulu-Natal have shown that receipt of disability grants did contribute significantly to securing the basic needs of AIDS-infected adults (Knight, Hosegood, & Timæus, 2013). In a similar vein, a quantitative analysis of households in the Free State province has found that receipt of different grants could contribute between 10% and 60% to the total poverty reduction in AIDS-affected households (Booyesen & van der Berg, 2005). *Post hoc t*-tests in the present analysis suggested that the disability grant might alleviate the impoverishing effects of AIDS illness. However, this finding must be interpreted cautiously as the analysis was based on cross-sectional data and recipient households could conceivably differ from non-recipient households in other, unmeasured aspects. Further research is required to explore associations with governmental grants (such as the disability grant, old age pension, and foster care grant) more thoroughly by using longitudinal data and controlling effectively for potential determinants of grant receipt/eligibility. Furthermore – although HIV/AIDS is potentially one of the criteria that translate into grant eligibility – only 10% of all AIDS-unwell caregivers in this sample receive the disability grant, suggesting that state-funded social security for AIDS-affected families remains deficient.

In contrast to prior studies (Dyson, 2003; Gillespie et al., 2007), AIDS illness was more prevalent in the rural than in the urban sample. It is conceivable that rates of AIDS sickness are higher in rural areas due to greater distance from clinics and health care centres and poor public transport. That is, there might be a shift in the HIV/AIDS epidemic moving from higher infection risks in urban centres to a higher risk for non-adherence and thus AIDS illness in rural communities. It is further conceivable that higher AIDS rates in rural areas are related to labour migrants who became infected with HIV/AIDS in urban centres and are returning home to be cared for (Clark, Collinson, Kahn, Drullinger, & Tollman, 2007; Lurie et al., 2003; Welaga et al., 2009). Yet, the present dataset did not include information on migrancy status of respondents and thus did not allow us to empirically examine the above arguments.

Several limitations of the current analysis are noteworthy. First, the cross-sectional study design makes it difficult to distinguish potential causes from effects and disentangle upstream and downstream effects as well as to assess the alleviating long-term impact of social security systems. Further work with longitudinal data is thus required.

The second limitation is the absence of information on the HIV status of household members other than the primary caregiver. For instance, because almost 90% of survey

respondents were female, detailed information on males – potentially primary breadwinners for some households – was lacking (see Masanjala, 2007; Poku, 2002; Russell, 2004). It is conceivable that the impoverishing effect of the disease becomes more pronounced when several household members suffer from HIV/AIDS, as the increased demand of caring would absorb more labour from income-generating activities. However, HIV/AIDS clusters within families and households (Poku, 2002; Whiteside, 2002) and it may be assumed that the primary caregiver's HIV status is somewhat indicative of a household's overall HIV-related vulnerability.

In summary, the findings of this study support evidence that suggests a positive relationship between poverty and HIV/AIDS. Although the magnitude of the correlation between poverty and AIDS illness was reduced when accounting for urban/rural residency, it should still be considered as 'socially important' (see Hemphill, 2003). With regard to the economic repercussions of AIDS illness, welfare grants have the potential to mitigate the detrimental downstream impact of the disease on poverty. Given that only 10% of households with AIDS-unwell caregivers in the current sample had access to the disability grant, it is important to consider means of enhancing access to welfare for those in need.

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## Disclosure statement

No potential conflict of interest was reported by the authors.

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